

#### Blackwell, Multi-calibration and Fairness

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## Blackwell Approachability

Definitions (k = 2 on the blackboard)

$$\mathcal{T}= ext{target set}$$
 $ec{U}(a,s)= ext{vector of utilities}\in R^k$ 
 $\overline{U}_T=\sum_{t=1}^T ec{U}(a_t,s_t)/T\in R^k$ 
 $c= ext{closest point to }\overline{U}_T$ 
 $d(u,\mathcal{T})= ext{distance from }u ext{ to the set }\mathcal{T}$ 

# Blackwell Approachability: Proof

$$\begin{array}{rcl} d(\overline{U}_{T+1},\mathcal{T}) & \leq & d(\overline{U}_{T+1},c) \\ (T+1)^2 d(\overline{U}_{T+1},\mathcal{T})^2 & \leq & (T+1)^2 d(\overline{U}_{T+1},c)^2 \end{array}$$

RHS = 
$$(T+1)^2 |\overline{U}_{T+1} - c|_2^2$$
  
=  $(T+1)^2 |\frac{T\overline{U}_T + U_{T+1}}{T+1} - c|_2^2$   
=  $|T(\overline{U}_T - c) + (U_{T+1} - c)|_2^2$   
=  $|T(\overline{U}_T - c)|^2 + |U_{T+1} - c)|_2^2 + \text{inner product}$   
 $\leq Td(\overline{U}_T, c)^2 + 4M^2$   
 $\leq 4(T+1)M^2$ 

$$d(\overline{U}_{T+1}) \leq 2M\sqrt{1/T} \rightarrow 0$$

#### Multi-calibration

- Goal: unbiased estimation of subgroups
  - called multi-calibration
  - Getting lots of attention in fairness
  - Gave a version of this talk at a week long symposium at Simons Foundation on multi-calibration

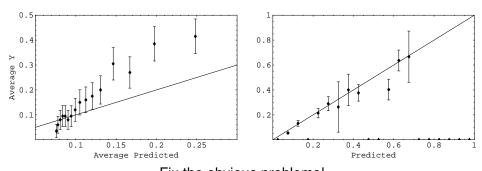
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  - horrible statistical properties
  - Many cells might even be empty
  - Can we only fix the k groups?

#### Multi-calibration

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  - horrible statistical properties
  - Many cells might even be empty
  - Can we only fix the k groups?
- We'll do it in an on line setting

# Statistics: Anything easily fixed isn't calibrated



Fix the obvious problems!

#### On-line Calibration

#### Calibration is a minimal condition for performance

- On sequence: 0 1 0 1 0 1 0 ...
- The constant forecast of .5 is calibrated
- The constant forecast of .6 is not calibrated
- The variable forecast of .1 .9 .1 .9 .1 .9 ... is not calibrated

#### On-line Calibration

#### Calibration is a minimal condition for performance

- On sequence: 0 1 0 1 0 1 0 ...
- The constant forecast of .5 is calibrated
- The constant forecast of .6 is not calibrated
- The variable forecast of .1 .9 .1 .9 .1 .9 ... is not calibrated
  - But the forecast .1 .9 .1 .9 ... is pretty good!
  - Yes, it has better "refinement."
  - But, it isn't calibrated.

#### Theorem

Blackwell approchability  $\Rightarrow$  no-internal regret  $\Rightarrow$  calibration.

#### Theorem

Calibrated forecasts exist.

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#### proof:

Apply mini-max theorem.

#### Theorem

Calibrated forecasts exist.

#### **Detailed proof:**

- Game between the statistician and Nature.
- Fine the value of a  $2^{2^T} \times 10^{2^T}$  matrix game.
- (Sergiu Hart: 1995 to 2023)

# These proofs were a bit fast

- These proofs are cute
- But still they take a few hours to understand

## These proofs were a bit fast

- These proofs are cute
- But still they take a few hours to understand
- So I doubt you got all the details
- I'll do a more useful proof
  - Uses least squares regression (so something you know)
  - Is practical (so details worth learning)
  - Solves the multi-calibration problem also

- We will falsify someone's claim by winning bets placed against them
- Claim:  $\hat{Y} \approx EY$ 
  - Prove it wrong by winning lots of money:

expected winnings = 
$$E\left(B\left(Y-\hat{Y}\right)\right)$$

- $(Y \hat{Y})$  is a "fair" bet
- B is amount bet

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- How to avoid being proven wrong by:

$$E\left(B\left(Y-\hat{Y}\right)\right)$$

(Start with bet B)

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- Claim:  $\hat{Y} \approx EY$ 
  - Prove it wrong by winning lots of money:

expected winnings = 
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- $(Y \hat{Y})$  is a "fair" bet
- B is amount bet
- How to avoid being proven wrong by:

$$Macau \equiv \max_{|B| \le 1} E\left(B\left(Y - \hat{Y}\right)\right)$$

(worry about worst bet)

- We will falsify someone's claim by winning bets placed against them
- Claim:  $\hat{Y} \approx EY$ 
  - Prove it wrong by winning lots of money:

expected winnings = 
$$E\left(B\left(Y-\hat{Y}\right)\right)$$

- $(Y \hat{Y})$  is a "fair" bet
- B is amount bet
- How to avoid being proven wrong by:

$$\min_{\hat{Y}} \max_{|B| \le 1} E\left(B\left(Y - \hat{Y}\right)\right)$$
(mini-max)

Y	$X_1$	$X_2$	$X_3$	$X_4$
<i>Y</i> <sub>1</sub>	X <sub>11</sub>	X <sub>12</sub>	X <sub>13</sub>	X <sub>14</sub>
Y <sub>2</sub>	$X_{21}$	$X_{22}$	$X_{23}$	X <sub>24</sub>
Y <sub>3</sub>	<i>X</i> <sub>31</sub>	$X_{32}$	$X_{33}$	X <sub>34</sub>
Y <sub>4</sub>	$X_{41}$	$X_{42}$	$X_{43}$	X <sub>44</sub>
:	÷	÷	÷	:
$Y_t$	$X_{t1}$	$X_{t2}$	$X_{t3}$	$X_{t4}$

Starting with our data that we observed up to time t

Y	$X_1$	$X_2$	$X_3$	$X_4$
<i>Y</i> <sub>1</sub>	X <sub>11</sub>	X <sub>12</sub>	X <sub>13</sub>	X <sub>14</sub>
<i>Y</i> <sub>2</sub>	X <sub>21</sub>	$X_{22}$	$X_{23}$	$X_{24}$
<i>Y</i> <sub>3</sub>	X <sub>31</sub>	$X_{32}$	<i>X</i> <sub>33</sub>	$X_{34}$
<i>Y</i> <sub>4</sub>	X <sub>41</sub>	$X_{42}$	$X_{43}$	$X_{44}$
:	:	:	:	:
$Y_t$	$X_{t1}$	$X_{t2}$	$X_{t3}$	$X_{t4}$

$$\hat{\beta}_t = \arg\min_{\beta} \sum_{i=1}^t (Y_i - \beta' X_i)^2$$

We can fit  $\hat{\beta}_t$  on everything up to time t

Y	$X_1$	$X_2$	<i>X</i> <sub>3</sub>	$X_4$			
<i>Y</i> <sub>1</sub>	X <sub>11</sub>	X <sub>12</sub>	X <sub>13</sub>	X <sub>14</sub>			
$Y_2$	<i>X</i> <sub>21</sub>	$X_{22}$	$X_{23}$	$X_{24}$			
<i>Y</i> <sub>3</sub>	<i>X</i> <sub>31</sub>	$X_{32}$	<i>X</i> <sub>33</sub>	$X_{34}$			
$Y_4$	$X_{41}$	$X_{42}$	$X_{43}$	$X_{44}$			
:	÷	:	÷	:			
$Y_t$	$X_{t1}$	$X_{t2}$	$X_{t3}$	$X_{t4}$			
	$X_{t+1,1}$	$X_{t+1,2}$	$X_{t+1,3}$	$X_{t+1,4}$	$\hat{eta}_t$	$\hat{Y}_{t+1}$	$=\hat{\beta}_t'X_t$

From a new  $X_{t+1}$  we can compute  $\hat{Y}_{t+1}$ 

Y	$X_1$	$X_2$	<i>X</i> <sub>3</sub>	$X_4$	$\hat{eta}$
<i>Y</i> <sub>1</sub>	X <sub>11</sub>	X <sub>12</sub>	X <sub>13</sub>	X <sub>14</sub>	0
Y <sub>2</sub>	X <sub>21</sub>	$X_{22}$	$X_{23}$	$X_{24}$	$\hat{\beta}_1$
Y <sub>3</sub>	<i>X</i> <sub>31</sub>	$X_{32}$	$X_{33}$	$X_{34}$	$\hat{\beta}_2$
Y <sub>4</sub>	X <sub>41</sub>	$X_{42}$	$X_{43}$	$X_{44}$	$\hat{\beta}_3$
:	:	:	÷	÷	:
$Y_t$	$X_{t1}$	$X_{t2}$	$X_{t3}$	$X_{t4}$	$\hat{\beta}_{t-1}$

Looking at only the first part of the data, we can generate:

$$\hat{\beta}_0$$
,  $\hat{\beta}_1$ ,  $\hat{\beta}_2$ ,  $\hat{\beta}_3$ ,  $\hat{\beta}_4$ , ...,  $\hat{\beta}_{t-1}$ 

Y	<i>X</i> <sub>1</sub>	$X_2$	<i>X</i> <sub>3</sub>	$X_4$	$\hat{eta}$	Ŷ
<i>Y</i> <sub>1</sub>	X <sub>11</sub>	X <sub>12</sub>	X <sub>13</sub>			$\hat{Y}_1 = 0$
Y <sub>2</sub>	<i>X</i> <sub>21</sub>	$X_{22}$	$X_{23}$	$X_{24}$	$\hat{\beta}_1$	$\hat{Y}_2 = \hat{\beta}_1' X_2$
Y <sub>3</sub>	<i>X</i> <sub>31</sub>	X <sub>32</sub>	<i>X</i> <sub>33</sub>	$X_{34}$	$\hat{\beta}_2$	$\hat{Y}_3 = \hat{eta}_2' X_3$
Y <sub>4</sub>	$X_{41}$	$X_{42}$	$X_{43}$	$X_{44}$	$\hat{\beta}_3$	$\hat{Y}_4 = \hat{eta}_3' X_4$
:	:	:		:	:	:
	$X_{t1}$	$X_{t2}$	$X_{t3}$	$X_{t4}$	$\hat{\beta}_{t-1}$	$\hat{Y}_t = \hat{\beta}'_{t-1} X_t$

Each of these leads to a next round

$$\hat{Y}_1, \quad \hat{Y}_2, \quad \hat{Y}_3, \quad \hat{Y}_4, \quad \dots, \quad \hat{Y}_t$$

Y	$X_1$	$X_2$	$X_3$	$X_4$	$\hat{eta}$	Ŷ
<i>Y</i> <sub>1</sub>	X <sub>11</sub>	X <sub>12</sub>	X <sub>13</sub>	X <sub>14</sub>	0	$\hat{Y}_1 = 0$
Y <sub>2</sub>	<i>X</i> <sub>21</sub>	$X_{22}$	$X_{23}$	$X_{24}$	$\hat{\beta}_1$	$\hat{Y}_2 = \hat{eta}_1' X_2$
Y <sub>3</sub>	<i>X</i> <sub>31</sub>	<i>X</i> <sub>32</sub>	$X_{33}$	$X_{34}$	$\hat{\beta}_2$	$\hat{Y}_3 = \hat{eta}_2' X_3$
Y <sub>4</sub>	$X_{41}$	$X_{42}$		$X_{44}$	$\hat{eta}_3$	$\hat{Y}_4 = \hat{eta}_3' X_4$
:	÷	:		:	:	:
$  Y_t  $			$X_{t3}$		$\hat{\beta}_{t-1}$	$\hat{Y}_t = \hat{\beta}'_{t-1} X_t$

#### Theorem (F. 1991, Forster 1999)

Such an on-line least squares forecast generates low regret:

$$\sum_{t=1}^{T} (Y_t - \hat{Y}_t)^2 - \min_{\beta} \sum_{t=1}^{T} (Y_t - \beta' X_t)^2 \leq O(\log(T))$$

Y	$X_1$	$X_2$	$X_3$	$X_4$	$\hat{eta}$	Ŷ
<i>Y</i> <sub>1</sub>	X <sub>11</sub>	X <sub>12</sub>	X <sub>13</sub>			$\hat{Y}_1 = 0$
Y <sub>2</sub>	<i>X</i> <sub>21</sub>	$X_{22}$	$X_{23}$	$X_{24}$	$\hat{eta}_1$	$\hat{Y}_2 = \hat{\beta}_1' X_2$
Y <sub>3</sub>	<i>X</i> <sub>31</sub>	<i>X</i> <sub>32</sub>	<i>X</i> <sub>33</sub>	$X_{34}$	$\hat{\beta}_2$	$\hat{Y}_3 = \hat{eta}_2' X_3$
Y <sub>4</sub>	$X_{41}$	$X_{42}$	$X_{43}$	$X_{44}$	$\hat{\beta}_3$	$\hat{Y}_4 = \hat{eta}_3' X_4$
:	÷	:		:	:	<u>:</u>
	$X_{t1}$	$X_{t2}$	$X_{t3}$	$X_{t4}$	$\hat{\beta}_{t-1}$	$\hat{Y}_t = \hat{\beta}'_{t-1} X_t$

Works no matter what the X's are.

Example: Use previous  $X_{t,i} = \hat{Y}_{t-i}$ . (F. and Stine 2021)

But we are going to go one better:  $X_t = \hat{Y}_t$ .

Y	$X_1$	$X_2$	<i>X</i> <sub>3</sub>	$X_4$	$\hat{eta}$	Ŷ
<i>Y</i> <sub>1</sub>	X <sub>11</sub>	X <sub>12</sub>	$\hat{Y}_1$	X <sub>14</sub>	0	$\hat{Y}_1 = 0$
Y <sub>2</sub>	<i>X</i> <sub>21</sub>	$X_{22}$	$\hat{Y}_2$	$X_{24}$	$\hat{\beta}_1$	$\hat{Y}_2 = \hat{\beta}_1' X_2$
Y <sub>3</sub>	<i>X</i> <sub>31</sub>	<i>X</i> <sub>32</sub>	$\hat{Y}_3$	<i>X</i> <sub>34</sub>	$\hat{\beta}_2$	$\hat{Y}_3 = \hat{\beta}_2' X_3$
Y <sub>4</sub>	$X_{41}$	$X_{42}$	$\hat{Y}_4$	$X_{44}$	$\hat{eta}_3$	$\hat{Y}_4 = \hat{eta}_3^{\prime\prime} X_4$
	÷	÷	÷	÷	1 :	:
$ Y_t $	$X_{t1}$	$X_{t2}$	$\hat{Y}_t$	$X_{t4}$	$\hat{\beta}_{t-1}$	$\hat{Y}_t = \hat{\beta}'_{t-1} X_t$

Theorem holds when one of the  $X_t$ 's is  $\hat{Y}_t$ !

Y	$X_1$	$X_2$	<i>X</i> <sub>3</sub>	$X_4$	$\hat{eta}$	Ŷ
<i>Y</i> <sub>1</sub>	X <sub>11</sub>	X <sub>12</sub>	Ŷ <sub>1</sub>	X <sub>14</sub>		$\hat{Y}_1 = 0$
Y <sub>2</sub>	<i>X</i> <sub>21</sub>	$X_{22}$	$\hat{Y}_2$	$X_{24}$	$\hat{\beta}_1$	$\hat{Y}_2 = \hat{\beta}_1' X_2$
Y <sub>3</sub>	<i>X</i> <sub>31</sub>	<i>X</i> <sub>32</sub>	$\hat{Y}_3$		$\hat{eta}_2$	$\hat{Y}_3 = \hat{\beta}_2' X_3$
Y <sub>4</sub>	$X_{41}$	$X_{42}$	$\hat{Y}_4$	$X_{44}$	$\hat{\beta}_3$	$\hat{Y}_4 = \hat{eta}_3' X_4$
	:	:	:	:	:	:
$ Y_t $		$X_{t2}$	^		$\hat{\beta}_{t-1}$	$\hat{Y}_t = \hat{\beta}'_{t-1} X_t$

#### Theorem ( ⇒ F. and Kakade 2008, F. and Hart 2018)

Adding the crazy calibration variable generates low macau:

$$(\forall i)$$
  $\sum_{t=1}^{T} X_{t,i}(Y_t - \hat{Y}_t) = O(\sqrt{T \log(T)})$ 

E(Y|X)Least squaresNormal equationsStatistics $\min_{\beta} \sum (Y_i - \beta \cdot X_i)^2$  $\sum X_i \ (Y_i - \beta \cdot X_i) = 0$ 

The normal equation is the same as:

$$\max_{\alpha} \sum_{i} \alpha' X_{i} (Y_{i} - \beta' X_{i})) = 0$$

Which is solved by the  $\beta$  minimizer:

$$\min_{\beta} \max_{\alpha} \sum_{i} \alpha' X_{i} (Y_{i} - \beta' X_{i})) = 0$$

E(Y X)	Least squares	Normal equations
Statistics	$\min_{\beta} \sum (Y_i - \beta \cdot X_i)^2$	$\min_{\beta} \max_{\alpha} \sum_{\alpha} \alpha \cdot X_i \ (Y_i - \beta \cdot X_i)$

E(Y X)	Least squares Normal equations	
Statistics	$\min_{\beta} \sum (Y_i - \beta \cdot X_i)^2$	$\left  \min_{\beta} \max_{\alpha} \sum_{\alpha} \alpha \cdot X_i \ (Y_i - \beta \cdot X_i) \right $
Probability	$\min_{f} E((Y - \underbrace{f(X)}_{aka})^{2})$	$(\forall g) \ E(g(X) \ (Y - f(X))) = 0$

The normal equation is the same as:

$$\max_{g} E\left(g(X)(Y - f(X))\right) = 0$$

Which is solved by the  $f(\cdot)$  minimizer:

$$\min_{f} \max_{g} E\left(g(X)(Y - f(X))\right) = 0$$

E(Y X)	Least squares	Normal equations
Statistics	$\min_{\beta} \sum (Y_i - \beta \cdot X_i)^2$	$\min_{eta} \max_{lpha} \sum_{lpha} lpha \cdot X_i \ (Y_i - eta \cdot X_i)$
Probability	$\min_{f} E((Y - \underbrace{f(X)}_{aka})^{2})$	$\min_{f} \max_{g} E(g(X) (Y - f(X)))$

E(Y X)	Least squares	Normal equations
Statistics	$\min_{eta} \sum (Y_i - eta \cdot X_i)^2$	$\min_{eta} \max_{lpha} \sum_{lpha} \alpha \cdot X_i \ (Y_i - eta \cdot X_i)$
Probability	$\min_{f} E((Y - \underbrace{f(X)}_{aka})^{2})$	$\min_{f} \max_{g} E(g(X) (Y - f(X)))$
online	low regret	low macau

Regret 
$$\equiv \sum_{t=1}^{T} (Y_t - \hat{Y}_t)^2 - \min_{\beta} \sum_{t=1}^{T} (Y_t - \beta \cdot X_t)^2$$

E(Y X)	Least squares	Normal equations
Statistics	$\min_{\beta} \sum (Y_i - \beta \cdot X_i)^2$	$\min_{eta} \max_{lpha} \sum_{lpha} \alpha \cdot X_i \ \left( Y_i - eta \cdot X_i  ight)$
Probability	$\min_{f} E((Y - \underbrace{f(X)}_{aka})^{2})$	$\min_{f} \max_{g} E(g(X) (Y - f(X)))$
online	low regret	low macau

$$\textit{Macau} \equiv \max_{\alpha: |\alpha| \le 1} \sum_{t=1}^{T} \alpha \cdot X_t \left( Y_t - \hat{Y}_t \right)$$

E(Y X)	Least squares	Normal equations
Statistics	$\min_{\beta} \sum (Y_i - \beta \cdot X_i)^2$	$\min_{\beta} \max_{\alpha} \sum_{\alpha} \alpha \cdot X_i \ (Y_i - \beta \cdot X_i)$
Probability	$\min_{f} E((Y - \underbrace{f(X)}_{aka})^{2})$	$\min_{f} \max_{g} E(g(X) (Y - f(X)))$
online	low regret	low macau

- ullet probability: Least squares  $\iff$  normal equations

E(Y X)	Least squares	Normal equations
Statistics	$\min_{\beta} \sum (Y_i - \beta \cdot X_i)^2$	$\min_{eta} \max_{lpha} \sum_{lpha} \alpha \cdot X_i \ \left( Y_i - eta \cdot X_i  ight)$
Probability	$\min_{f} E((Y - \underbrace{f(X)}_{aka})^{2})$	$\min_{f} \max_{g} E(g(X) (Y - f(X)))$
online	low regret	low macau

#### Take Aways

on-line low regret 

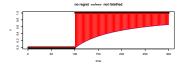
⇔ on-line low macau

## 

#### No regret ⇒ not falsified

								T+2		
$Y_t$	0	0	0	0	 0	1	1	1	1	 1
$X_t$	1	1	1	1	 1	1	1	1	1	 1
$\hat{Y}_t$	0	0	0	0	 0	0	1	$\frac{2}{T+1}$	3 7±2	 2 3

#### How about a bet?

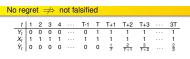


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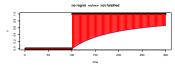
t	1	2	3	4	 Т	T+1	
$Y_t$	0	1	0	1	 0	1	
$X_t$	1	1	1	1	 1	1	
$\hat{Y}_t$	.6	.4	.6	.4	 .6	1 1 .4	

- Macau is zero
- Regret is T/9
- So: low macau ⇒ low regret

## 



#### How about a bet?



#### Not falsified ⇒ no regret

t	1	2	3	4	 Т	T+1	
$Y_t$	0	1	0	1	 0	1	
$X_t$	1	1	1	1	 1	1 1 .4	
Ŷ,	.6	.4	.6	.4	 .6	.4	

- Macau is zero
- Regret is T/9
- So: low macau ⇒ low regret

(Skipping these proofs)

$$C(a) = \sum_{t=1}^{T} c_t(a)$$
  $a^* \equiv \arg\min_{a} C(a)$ 

- Supposed each  $c_t(\cdot)$  is convex
- Goal: play a to minimize C(a)
- Eg: We could use SGD on  $\nabla c_t()$
- called "on-line convex optimization" with regret:

regret 
$$\equiv \sum_{t=1}^{T} (c_t(\hat{a}_t) - c_t(a^*))$$

$$C(a) = \sum_{t=1}^{T} c_t(a)$$
  $a^* \equiv \arg\min_{a} C(a)$ 

The regret is bounded by the gradient:

regret 
$$=\sum_{t=1}^{T}(c_t(\hat{a}_t)-c_t(a^*))$$
  
 $\leq \sum_{t=1}^{T}(\hat{a}_t-a^*)\cdot \nabla c_t(\hat{a}_t)$ 

$$C(a) = \sum_{t=1}^{T} c_t(a)$$
  $a^* \equiv \arg\min_{a} C(a)$ 

The regret is bounded by the gradient:

$$\begin{split} \text{regret} & = \sum_{t=1}^T (c_t(\hat{a}_t) - c_t(a^*)) \\ & \leq \sum_{t=1}^T (\hat{a}_t - a^*) \cdot \nabla c_t(\hat{a}_t) \\ & = \sum_{t=1}^T (\hat{a}_t - a^*) \cdot \left( \nabla c_t(\hat{a}_t) - \widehat{\nabla c_t}(\hat{a}_t) \right) + (\hat{a}_t - a^*) \cdot \widehat{\nabla c_t}(\hat{a}_t) \end{split}$$

$$C(a) = \sum_{t=1}^{T} c_t(a)$$
  $a^* \equiv \arg\min_{a} C(a)$ 

The regret is bounded by the gradient:

regret 
$$= \sum_{t=1}^{T} (c_t(\hat{a}_t) - c_t(a^*))$$

$$\leq \sum_{t=1}^{T} (\hat{a}_t - a^*) \cdot \nabla c_t(\hat{a}_t)$$

$$= \sum_{t=1}^{T} (\hat{a}_t - a^*) \cdot \left( \nabla c_t(\hat{a}_t) - \widehat{\nabla c_t}(\hat{a}_t) \right) + (\hat{a}_t - a^*) \cdot \widehat{\nabla c_t}(\hat{a}_t)$$

$$(macaul) \qquad (zero @ \hat{a}_t)$$

$$C(a) = \sum_{t=1}^{T} c_t(a)$$
  $a^* \equiv \arg\min_{a} C(a)$ 

The regret is bounded by the gradient:

$$\begin{aligned} \text{regret} &= \sum_{t=1}^{T} (c_t(\hat{a}_t) - c_t(a^*)) \\ &\leq \sum_{t=1}^{T} (\hat{a}_t - a^*) \cdot \nabla c_t(\hat{a}_t) \\ &= \sum_{t=1}^{T} (\hat{a}_t - a^*) \cdot \left( \nabla c_t(\hat{a}_t) - \widehat{\nabla c_t}(\hat{a}_t) \right) + (\hat{a}_t - a^*) \cdot \widehat{\nabla c_t}(\hat{a}_t) \end{aligned}$$

regret ≤ macau

#### Calibration Theorem

### Theorem ( $\implies$ F. and Kakade 2008, $\iff$ new)

Let R be the quadratic regret of a forecast  $\hat{Y}_t$  against a linear regression on  $X_t$ . Let M be the Macau of  $\hat{Y}_t$  using linear functions of  $X_t$  to create falsifying bets. Then if we have the crazy calibration variable (i.e.  $[X_t]_0 = \hat{Y}_t$ ), then

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Proof sketch: Consider the forecasts  $(1 - w)\hat{Y}_t + w\alpha \cdot X_t$  for the any  $\alpha$ . Let Q(w) be the total quadratic error of this family of forecast. The following are equivalent:

- $Q(0) \leq Q(w)$  (No regret condition)
- Q'(0) is zero. (No macau condition)

### Calibration Theorem

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Note: Typically,  $R = O(\log(T))$  iff  $M = \tilde{O}(\sqrt{T})$  for the actual algorithms I know.

(S. Rakhlin and D. Foster have a proof for IID.)

# Recipe for good decisions

- List bets that you would make to show  $\hat{a}_t$  is not optimal
- Convert these to regression variables
- Add the crazy-calibration variable
- Run a low regret least squares algorithm
- Make decision based on this forecast

## That is Macau

## Take Aways

crazy-Calibration + low-regret  $\iff$  low-macau  $\implies$  good decisions

## Fairness and incentives

- Consider predicts used for college admissions
  - We'll call the prediction: SAT
  - We'll call the Y variable: GPA
- We are interested in fair incentives
  - The incentive story works better for employment,
  - But the names will be useful, so we'll stick with college admissions

## Regress *Y* on *X* or regression *X* on *Y*?

Basic discrimination:

$$E(GPA|blue, SAT=x) > E(GPA|orange, SAT=x)$$

- Better off being orange
- Richard Posner argued economics would drive it out
- So it simply doesn't exist due to "rationality"

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- But even if

$$E(GPA|blue, SAT=x) = E(GPA|orange, SAT=x)$$

we might have:

$$E(SAT|blue, skill=y) < E(SAT|orange, skill=y)$$

So still better off being Orange!

# Backwards regression

Traditional regression:

$$\min_{f} E\left((Y - f(X))^{2}\right)$$

Reverse regression:

$$\min_{g} E\left((g(Y) - X)^{2}\right)$$

- Even if f() and g() are linear,  $f \neq g^{-1}$
- (unless we have a perfect fit)
- Called regression to the mean

## No measurement of skill

- We don't have skill, but we do have GPA
- So, regress SATs on GPAs and make that calibrated
  - Fair incentives
  - Economics won't come to this solution with Laissez-faire
  - Needs government intervention (F. and Vohra, 1992)

## No measurement of skill

- We don't have skill, but we do have GPA
- So, regress SATs on GPAs and make that calibrated
  - Fair incentives
  - Economics won't come to this solution with Laissez-faire
  - Needs government intervention (F. and Vohra, 1992)
- Fairness then is best approximated by:

$$E(SAT|blue, GPA=y) \approx E(SAT|orange, GPA=y)$$

## References: Three different Fosters

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Take Aways

crazy-Calibration + low-regret  $\iff$  low-macau

2: Accuracy is not the same as fairness

Take Aways

crazy-Calibration + low-regret ← low-macau

2: Accuracy is not the same as fairness

Thanks!